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Original Article

Simultaneous Optimization of Multiple Response Variables for the Gelatin-chitosan Microcapsules Containing Angelica Essential Oil

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Abstract

Angelica essential oil (AO), a major pharmacologically active component of Angelica sinensis (Oliv.) Diels, possesses hemogenesis, analgesic activities, and sedative effect. The application of AO in pharmaceutical systems had been limited because of its low oxidative stability. The AO-loaded gelatin-chitosan microcapsules with prevention from oxidation were developed and optimized using response surface methodology. The effects of formulation variables (pH at complex coacervation, gelatin concentration, and core/wall ratio) on multiple response variables (yield, encapsulation efficiency, antioxidation rate, percent of drug released in 1 h, and time to 85% drug release) were systemically investigated. A desirability function that combined these five response variables was constructed. All response variables investigated were found to be highly dependent on the formulation variables, with strong interactions observed between the formulation variables. It was found that optimum overall desirability of AO microcapsules could be obtained at pH 6.20, gelatin concentration 25.00%, and core/ wall ratio 40.40%. The experimental values of the response variables highly agreed with the predicted values. The antioxidation rate of optimum formulation was approximately 8 times higher than that of AO. The in-vitro drug release from microcapsules was followed Higuchi model with super case-II transport mechanism.

Keywords: microcapsules; angelica essential oil; multiresponse optimization; gelatin; sustained release; antioxidation rate.

Introduction

Angelica essential oil (AO), a major pharmacologically active component of Angelica sinensis (Oliv.) Diels, possesses hemogenesis, analgesic activities and sedative effect, and finds application in the treatment of a range of conditions including menstrual disturbance and anemia (1). Over 40 compounds have been identified in AO (2) with the major constituent

being ligustilide (3). Numerous studies have shown that ligustilide has a neuroprotective effect against ischaemia-reperfusion injury via anti-apoptosis and anti-oxidation in neurons (4-6).

application of AO in However, the pharmaceutical systems has been limited because of its low oxidative stability (7). When AO is processed, used, and stored as a drug or dietary supplement, accelerated degradation of AO can result in decreased drug effect and off-flavor development (8). In addition, the above benefits can be effectively acquired when AO is taken

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repeatedly and continuously for a given period of time, because of the short retention time of AO in plasma. Thus, the sustained delivery of AO and prevention of AO from oxidation were two challenges facing the oral delivery systems design works.

Microencapsulation technology was considered as a promising technology to protect from oxidation and maintain sustained release for AO. Complex coacervation, caused by electrostatic attraction of two oppositely charged colloids, has been used widely to microencapsulate drug components (9, 10). The coating materials can be prepared from gums, proteins, polysaccharides, lipids and synthetic polymers (11, 12). However, for a medicinal product, a bio-polymeric system based on natural polysaccharides or protein could provide a safe, inert and effective oral delivery matrix (13, 14). Gelatin B has excellent emulsibility, high hydrophilicity, and particular gelling property (15). Chitosan has been reported to be very suitable for preparation of microcapsules for controlled drug release. The pharmaceutical and food applications of chitosan have remarkably increased over recent years because of its low production costs, biodegradability, biocompatibility and recent FDA approval (16). In this work, the coacervate of gelatin and chitosan was selected as the coating materials.

A problem facing the product development community is the selection of a set of conditions which will result in a product with a desirable combination of properties. Response surface methodology (RSM) is useful in improvement and optimization of pharmaceutical formulations by finding the analytical relationship between input and output variables considered in experiments. RSM also has the ability to produce an approximate function using a smaller amount of data and fewer numbers of experiments. However, most previous applications based on RSM have only dealt with a single-response problem (17, 18) and multi-response problems have received only limited attention.

During optimization of gelatin-chitosan microcapsules (GCM), usually several response variables were to be optimized. In many cases, these responses were competing, i.e., improving one response might have an opposite effect on

another one, which further complicated the situation. A method to resolving the problem involving the simultaneous optimization of multiple response variables is through the use of a desirability function that combined all the responses into one measurement (19, 20). Studies have showed that a desirability function method is useful in simultaneously optimizing several response variables (21, 22).

In this paper, the effect of formulation variables of GCM containing AO on several response variables were investigated, and subsequently, a formulation with excellent antioxidation and sustained-release properties was obtained using RSM after constructing a desirability function that combined all response variables.

Materials and Methods Materials

AO was obtained by supercritical carbon dioxide extraction method and donated by Humei Natual Spices Oil Refineries Factory. The natural proportion of ligustilide in AO was 52%. Gelatin from porcine skin, type B, Bloom 260, was obtained from Xiamen Huaxuan Gelatin Co., Ltd. Chitosan, low viscosity (22 cps), 200 thousand Daltons molecular weight, 88.1% deacetylation, was purchased from Shandong AK Biotech Ltd. Glutaraldehyde (50%), as a crosslinker, was provided by Kermel Co., Ltd. The other chemicals and solvents used in this work were of analytical grade, purchased from Sigma-Aldrich Co., Ltd. Deionised water (electrical conductivity $< 2 \mu S cm^{-1}$) was used throughout all the experiments.

Preparation of GCM

GCM were prepared using the complex coacervation method. Gelatin solution (concentrations: 6.34-23.66%, w/w) was prepared by swelling gelatin in deionised water followed by heating (50 °C) until the appearance of a clear solution. Chitosan was dissolved in 1% (w/v) acetic acid solution by stirring overnight until a clear solution was obtained. The concentration ratio of gelatin to chitosan was 10/1 described by Silva *et al.* (23). O/W emulsion was prepared by adding AO (core/wall ratios: 19.02-70.98%) into the above gelatin solution at high shear (Fluko

Homogenizers, model FA25, USA) rate of 10000 rpm (30-40 s), and diluted 2 times with deionised water. Then chitosan solution was added to the emulsion under stirring rate of 400 rpm for 30 min. Initially, the coacervation of chitosan and gelatin was brought about by gradual addition of 2% sodium hydroxide solution. Then the pH of mixture was adjusted to 5.32-6.18. In this stage, GCM were formed. This mixture was cooled to 20 °C in water-bath for obtaining coacervate precipitation. The cross-linking of the GCM was achieved by addition of a certain amount of 5% glutaraldehyde solution (0.3-0.5 g glutaraldehyde / 1 g wall materials). After 20 min, the mixture was heated and maintained at 30-40 °C, and stirred for 40 min. This mixture was then cooled to room temperature. The microcapsules were filtered by vacuum pump, then washed with deionised water at 35 °C to remove excess glutaraldehyde, and then dried.

Determination of yield

Yield (%, w/w) was calculated as follows: Yield (%, w/w) = $W_d/W_s \times 100$ (eq. 1), where W_d is the weight of dried microcapsules recovered, W_s is the total weight of the wall materials and AO initially added during the batch preparation.

AO assay in microcapsules

In AO, the relative amount of ligustilide was constant. Therefore, the content of AO in the microcapsules was calculated according to the following formula: $C_A = C_L/P_L \times 100$ (eq. 2), where C_A is the content of AO, C_L is the content of ligustilide, P_L is the natural proportion of ligustilide in AO (i.e., 52%).

High performance liquid chromatography (HPLC) method was established for the determination of ligustilide. A series of known concentrations in the range 1.25-100 μ g mL⁻¹ of ligustilide in mobile phase containing 75% acetonitrile and 25% deionised water were determined at the detective wavelength 326 nm (Shimadzu, model LC-10AT, Japan). A C₁₈ column (250 mm × 4.6 mm, 5 μ M, Dikma Technologies, China) was used at room temperature, and flow rate was 0.80 mL min⁻¹. The respective peak areas were recorded and plotted.

A certain amount of GCM was accurately

weighed and dispersed in a known volume of mobile phase. After staying overnight, this dispersion was filtrated, and the content of ligustilide in the resultant filtrate was measured by HPLC method as described above. Based on the resulting value, the content of AO in the microcapsules was calculated. Each experiment was carried out in triplicate.

Measurement of encapsulation efficiency (EE)

The EE of the microcapsules was determined as follows: EE (%) = $C_a/C_i \times 100$ (eq. 3), where C_a is the actual drug content in the microcapsules, C_i is the content of the drug initially added during the batch preparation.

Stability of microcapsules against oxidation

Oxidation is the major degradation pathway of AO (7). The stability of microcapsules against oxidation was studied in a 9-day acceleration test (60 °C, 75% relative humidity, exposed to air) (24, 25). The samples were collected at regular intervals. AO assay in GCM was determinated at the different time points. For unencapsulated AO, the contents of ligustilide at the various intervals were determined by HPLC as described in "AO assay in microcapsules", and the contents of unoxidated AO were calculated. Antioxidation rate (AR) was determined by substituting the resulting values in the following expression: AR (%) = $C_n/C_0 \times 100$ (eq. 4), where C_n and C_0 are the unoxidated AO content in microcapsules or in unencapsulated AO at Day n and Day 0, respectively.

Release of AO from GCM in-vitro

The release of AO from the GCM was investigated using a dissolution tester (Huanghai, model ZRS-8G, China) under a constant oar speed of 50 rpm at 37 °C. The phosphate buffer at pH 7.4 was used as the dissolution media (1000 mL). At appropriate time intervals, 5 mL samples were withdrawn and filtered through a 0.45 μ M Millipore membrane filter.

The ligustilide concentration in the dissolution media was determined by HPLC as described in "AO assay in microcapsules", and the concentration of AO was calculated. For unencapsulated AO, the percent of released drug

was calculated based on the added amount of AO. For GCM, the percent was obtained based on the results of AO assay in GCM, Then, the percent of released drug was plotted vs. time. The percent of drug released in 1 h (P_1) and time to 85% drug release (t_{85}) were calculated.

To evaluate the kinetics of drug release from the microcapsules, *in-vitro* release data were analyzed according to zero-order (eq, 5), firstorder (eq, 6), Higuchi (eq, 7), and Korsmeyer-Peppas (eq, 8) equations:

$$M/M_{\infty} = 100 (1-k_0 t)$$
 (eq. 5)

$$M/M_{x} = 100 (1 - \exp(k_{1}t))$$
 (eq. 6)

$$M/M_{\odot} = k_{b} t^{0.5}$$
 (eq. 7)

$$M/M = kt^n$$
 (eq. 8)

where M/M_{∞} is the fractional release of the drug in time t, n is the release exponent, indicative of the transport mechanism, k_0 , k_1 , k_h and k are constants incorporating geometrical and structural characteristics of the macromolecular network system and the drug (26, 27). The release exponent for polymeric controlled delivery systems of spherical geometry has values of $n \le 0.43$ for Fickian release (diffusion-controlled release), $0.43 < n \le 0.85$ for non-Fickian release (anomalous transport) and n > 0.85 for super case-II transport (relaxation-controlled release) (28). All data were analyzed using OriginPro 8 (OriginLab, UK).

Scanning electron microscopy (SEM) of GCM

The outer structures of the microcapsules were studied by SEM. Dried microcapsules were mounted on metal stubs and coated with gold (20 nm thickness) using an ion coater (Eiko Engi-neering, model IB-2, Japan). Accelerating voltages of 5 kV was used to observe the morphologies of the gold coated microcapsules. The samples were determined by image processing software (Image J, NIST) and captured by automatic image-capturing software.

Overall desirability (OD) function

To combine the five measured responses in one OD function, individual desirability

functions have to be calculated first. Individual desirability function involves transformation of each estimated response variable Y_i to a desirability value d_i , where $0 \le d_i \le 1$. The value of d_i increases as the "desirability" of the corresponding response increases (29). In this work, two methods were used in the calculation of individual desirability functions.

For response variables that were desired to be maximized, such as yield (w/w, %), EE (%), and AR (%) of microcapsules, d_i could be calculated as follows (19,20):

$$\mathbf{d}_i = \begin{cases} 0 & f \quad Y_i \leq Y_{i-\min} \\ \left[\frac{Y_i - Y_{i-\min}}{Y_{i-\max} - Y_{i-\min}}\right]^r & f \quad Y_{i-\min} < Y_i < Y_{i-\max} \\ 1 & f \quad Y_i \geq Y_{i-\max} \end{cases}$$

where Y_i are the actual observed response values of type i. Y_{i-min} and Y_{i-max} are minimum and maximum acceptable values of response i, respectively. r is a positive constant and is known as weight. For these three response variables, the more general linear-scale desirability function was used, i.e., r = 1 (30). Table 1.

showed the minimum and maximum acceptable values for these three response variables.

For P_1 and t_{85} , the two selected parameters of sustained-release characteristics, we used Harrington's exponential function (31), which was described as follows: based on the distribution of values, a desirability value (*d*) of 0.4 was assigned to P1 of 8%, and a value of 0.8 was assigned to p1 of 12% in the desirability scale with a maximum of 1.0. Similarly, t_{85} of 5 h was given a desirability value (*d*) of 0.2, and a t_{85} of 12 h was assigned a value of 0.8. Each of these two desirability values (*d*) was transformed to a dimensionless response (*Y'*) using the equation:

$$Y' = -[\ln(-\ln d)]$$
 (eq. 10)

From the two paired values of Y and Y', the following linear transformation equation was calculated.

$$Y' = b_0 + b_1 Y$$
 (eq. 11)

Table 1. Minimum and maximum acceptable levels of the response variables.

Response Variables	Minimum Level (%)	Maximum Level (%)
Yield	65	85
EE	75	95
AR	50	80

EE = encapsulation efficiency; AR = antioxidation rate.

Where b_0 and b_1 are constants, and in this case they were found to be equal to -2.379 and +0.353 for P_1 , and -2.13 and +0.33 for t_{85} , respectively. The desirability value of each formulation was calculated from the Y' value using the exponential equation:

$$D = \exp\{-[\exp[(-Y')]\}\$$
 (eq. 12)

The individual desirabilities were then combined using the geometric mean.

$$OD = (d_1 \times d_2 \times ... \times d_k)^{1/k}$$
 (eq. 13)

This single value of OD gives the overall assessment of the desirability of the combined response levels.

Experimental design and statistical analysis

The effects of three independent variables, namely pH at complex coacervation (X₁), gelatin concentration (X2, w/w, %), core/wall ratio (X3, %), on the OD of multiple response variables, including yield (w/w, %), EE (%), AR (%), and sustained-release profile (P_1, t_{85}) were studied using a three factor central composite design (CCD). Twenty microcapsules samples were established based on the CCD with three independent variables at five levels on each variable. The center point was repeated six times to calculate the reproducibility of the method. Multiple regression analysis was applied for prediction of the linear, quadratic and interaction terms of the independent variables in the RSM. Regression analysis was performed to estimate the response function as polynomial model:

$$Y = \beta_0 + \sum \beta_i x_i + \sum \beta_{ii} x_i^2 + \sum \beta_{ii} x_i x_i$$
 (eq. 14)

Where *Y* is response calculated by the model,

 β_0 is a constant, and β_i , β_{ii} , and β_{ij} are linear, squared and interaction coefficients, respectively (17). Data were modeled by multiple regression analysis. The significant terms in the model were found by analysis of variance for each response. The statistically significant parameters at the 95% significance level were only selected for the model construction. Experimental data were compared with the fitted values predicted by the models in order to verify the adequacy of the regression models. Each experiment was repeated in triplicate.

Results and Discussion

Optimization of GCM by RSM

Table 2. listed the evaluation results of each formulation in terms of the individual responses, along with the calculated OD function values. The results of the applied statistical tests indicated that five response variables measured in this study as well as the OD function showed good fitting to different models. The fitting equations after models were simplified, the corresponding r² values of models, and significance levels of individual parameters were listed in Table 3.

Yield (%)

From the fitting equation (Table 3) and contour plots (Figure. 1a & 1b), yield increased in a linear manner. Gelatin concentration and core/wall ratio had greater influence on yield than pH. The positive coefficient of X_1 referred to the increase in yield with increasing pH in the experimental range. The results of plot experiments indicated that insufficient reaction between gelatin and chitosan would occur at Ph < 5.30, and pH > 6.30 would bring about the low solubility of chitosan. A narrow range of pH (5.32-6.18) for complex coacervation were

Table 2. Experimental levels of the independent variables and values of the five measured responses and OD function for the prepared GCM formulations

Run – Number	Independent variables			Response variables					
	рН (X ₁)	gelatin conc. (X ₂ , %)	core/wall ratio (X ₃ ,%)	Yield (%)	EE (%)	AR (%)	P ₁ (%)	t ₈₅ (h)	OD
1	5.75	6.34	45.00	67.06	86.76	62.74	12.70	4.50	0.27
2	5.75	23.66	45.00	87.58	91.13	80.76	8.37	14.50	0.76
3	5.50	20.00	30.00	78.71	83.86	74.15	9.75	11.00	0.50
4	5.75	15.00	19.02	64.00	83.50	64.65	12.87	5.00	0
5	6.18	15.00	45.00	80.61	97.00	66.38	10.66	13.00	0.74
6	5.75	15.00	45.00	78.86	89.01	73.42	9.74	11.50	0.63
7	5.75	15.00	45.00	75.40	90.38	73.00	8.90	12.00	0.60
8	5.75	15.00	45.00	74.97	91.45	73.74	11.03	12.50	0.66
9	6.00	20.00	60.00	87.37	94.69	81.23	6.05	13.50	0.70
10	5.75	15.00	45.00	79.60	90.81	73.21	9.57	11.50	0.66
11	5.32	15.00	45.00	71.14	86.29	74.24	8.99	9.50	0.48
12	6.00	10.00	60.00	78.86	94.06	70.13	9.40	7.50	0.60
13	6.00	10.00	30.00	69.14	90.61	54.61	13.24	5.50	0
14	6.00	20.00	30.00	78.67	90.60	70.88	10.98	12.50	0.66
15	5.50	10.00	30.00	63.94	81.56	67.80	13.31	5.00	0
16	5.50	10.00	60.00	79.57	90.19	75.29	8.41	5.00	0.48
17	5.75	15.00	70.98	83.49	93.99	77.67	6.67	6.50	0.57
18	5.75	15.00	45.00	75.17	91.33	72.39	8.95	12.50	0.62
19	5.50	20.00	60.00	82.67	91.51	73.96	4.85	10.00	0.50
20	5.75	15.00	45.00	74.93	89.81	74.31	10.29	12.00	0.62

GCM = gelatin-chitosan microcapsules; conc. = concentration; EE = encapsulation efficiency; AR = antioxidation rate; P_1 = % of drug released in 1 h; t_{85} = time to 85% drug release; OD = overall desirability.

selected in this experimental design. In this range, pH had a weaker influence on yield. The yield was also increased by the gelatin concentration increase. Gelatin played an important role in the formation of a stable emulsion because of itself's emulsification. Studies indicated that beads formed from low concentrations of polymer in solution were expected to be weaker. Random collision and contact of spheres caused blending of weaker particles into each other and resulted in non-uniform shape (32).

The irregularly-shaped droplets that were formed in the absence or at extremely low concentrations of gelatin could stick to the stirrer, or to one another, forming big lumps, which could reduce the yield, as demonstrated by Chung *et al.* (33). In our previous experiments, the 15-25% concentration of gelatin was enough to precipitate-GCM. When the concentration was higher than 25%, it was difficult to gain product due to high viscosity and heavy gelatinization below 25 °C. Increasing core/wall ratio increased the yield.

This increase corresponded to the positive coefficients of X_3 . A higher core/wall ratio resulted in less sticky droplets, which reduced the loss caused by droplets aggregation in the post-processing step (34).

Table 3. The fitting models, equations, and statistical parameters of five measured responses and OD function.

Response Variables	Models	Equations ^a	\mathbf{r}^2	Pb	Lack of fit
Yield (%)	Linear	Yield (%) = $+76.59 + 1.83X_1 + 5.10X_2 + 5.13X_3$	0.8660	≤0.0001	NSS
EE (%)	Quadratic	EE (%) = $+90.47+2.96X_1+0.84X_2+3.00X_3-1.09X_1X_3-0.54X_2^2-0.60X_3^2$	0.9548	≤0.0113	NSS
AR (%)	Quadratic	AR (%) = +73.35-2.00 X_1 +4.54 X_2 +3.98 X_3 +2.79 X_1 X_2 +2. 32 X_1 X_3 -1.61 X_2 X_3 -1.04 X_1 ² -0.54 X_2 ² -0.74 X_3 ²	0.9769	≤0.0031	NSS
P1 (%)	Linear	P_1 (%) = +9.74+0.45 X_1 -1.45 X_2 -2.09 X_3	0.9153	≤0.0001	NSS
t85 (h)	Quadratic	t_{85} (h) = +12.00+1.01 X_1 +2.95 X_2 +0.33 X_3 +0.50 X_1X_3 - 0.27 X_1^2 -0.85 X_2^2 -2.10 X_3^2	0.9896	≤0.0001	NSS
OD	Quadratic	OD = $+0.63+0.067X_1+0.15X_2+0.15X_3+0.03 X_1X_2-0.13$ $X_2X_3-0.047 X_2^2-0.12 X_3^2$	0.9780	≤0.0001	NSS

EE = encapsulation efficiency; AR = antioxidation rate; P_1 = % of drug released in 1 h; t_{85} = time to 85% drug release; OD = overall desirability; NSS = not statistically significant.

EE (%)

EE showed good fitting to the quadratic model (Table 3). The coefficients of X_1 , X_2 , and X₃ in the fitting equation and the contour plots (Figure. 1c & 1d) indicated that the effects of pH at complex coacervation and core/wall ratio on EE were greater than gelatin concentration. EE increased as pH was increased (positive coefficient of X₁). This could be attributed to the higher degree of complex coacervation reaction between gelatin and chitosan at higher pH (35, 36). As the core/wall ratio increased, EE significantly increased (positive coefficient of X_2). A possible explanation for this was that only a fixed amount of AO was probably lost to the external phase during the formation of microcapsules and this loss obviously had a more detrimental effect on EE with relatively lower drug content (37). After reaching a maximum level, EE started to decline slightly as the core/wall ratio was further increased. This was indicated by the negative coefficient of the square term of X₂ and could be attributed to the practical limitation to the amount of drug that can be incorporated in the microcapsules (38, 39). Within the experimental range, gelatin concentration had the smallest effect on EE; this effect, however, was statistically significant.

Increasing gelatin concentration increased EE. The only statistically significant interaction in the EE equation occurred between pH at complex coacervation and core/wall ratio. This interaction was of a negative magnitude.

AR (%)

The compact degree of surface was an important factor that influenced the AR and sustained-release quality of microcapsules. Studies indicated that the compact surface (Figure. 2) could help to protect core material from oxygen, heat, and so on (40). The AR of microcapsules was significantly influenced by all three variables. This was manifested by square, linear, and interaction terms in the fitting equation (Table 3). AR showed good fitting to the quadratic model. The negative coefficients of X₁ and X₁² indicated that AR increased as pH at complex coacervation decreased. This could be attributed to the less compact surface of microcapsules at higher pH (41). Based on the fitting equation, AR was significantly dependent on the gelatin concentration and core/wall ratio. Initially, as the gelatin concentration or core/wall ratio increased, the surface of microcapsules was more compact (42) and AR increased (positive coefficients of X₂ and X₃). However, at higher

^a Where X₁, X₂, and X₃ represent the independent variables: pH at complex coacervation, gelatin concentration (w/w, %), core/wall ratio (%), respectively.

^b Significance levels of the individual parameters.

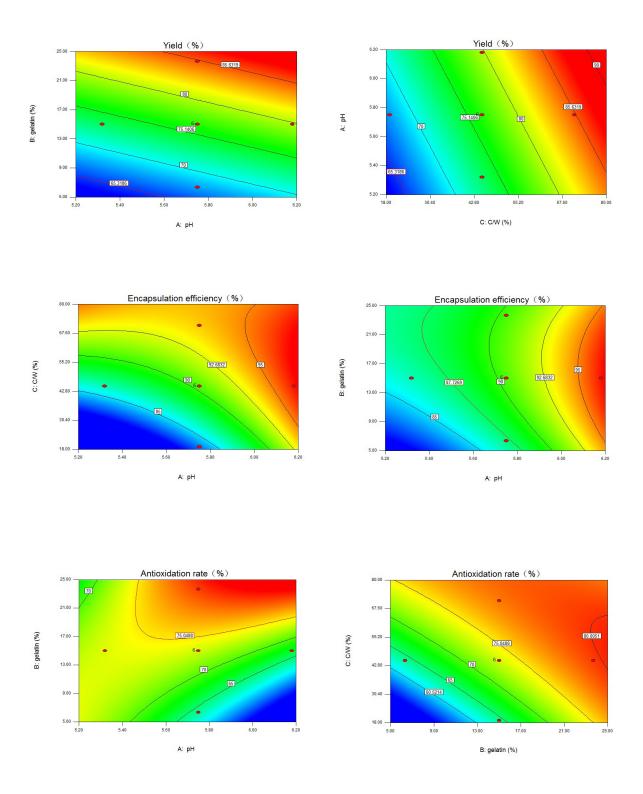


Figure 1. Contour plots of three response variables. a) Contour plot of yield, gelatin concentration vs. pH with core/wall ratio = 45%; b) Contour plot of yield, pH vs. core/wall ratio with gelatin concentration = 15%; c) Contour plot of encapsulation efficiency, pH vs. core/wall ratio with gelatin concentration = 15%; d) Contour plot of encapsulation efficiency, gelatin concentration vs. pH with core/wall ratio = 45%; e) Contour plot of antioxidation rate, gelatin concentration vs. pH with core/wall ratio = 45%; f) Contour plot of antioxidation rate, core/wall ratio vs. gelatin concentration with pH = 5.75. C/W = core/wall ratio.

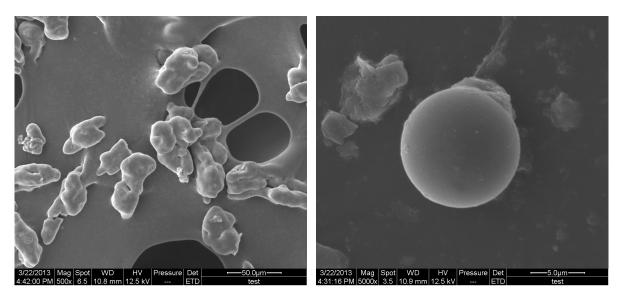


Figure 2. Scanning electron micrograph of gelatin-chitosan microcapsules. a) magnification of 500x; b) magnification of 5000x.

gelatin concentrations or at higher core/wall ratios, AR decreased as gelatin concentration or core/wall ratio increased (negative coefficients of X_2^2 and X_3^2). This could be related to parallel effects of these factors on EE.

The AR plots of AO and GCM were shown in Fig. 3. After 9-day acceleration tests, the drug content of GCM was 81.77%, which was much higher than that of AO. The AR of GCM was approximately 8 times higher than one of AO.

Sustained-release profile (P_1, t_{85})

Different release parameters had been used as response variables to describe and optimize the release process. The most frequently used parameters included the percent of drug released at certain time point, the time needed to release a certain percent of the drug, dissolution curve shape factor, dissolution rate constant, and the release order (43). In this study, P_1 and t_{85} were selected as the response variables to ensure full

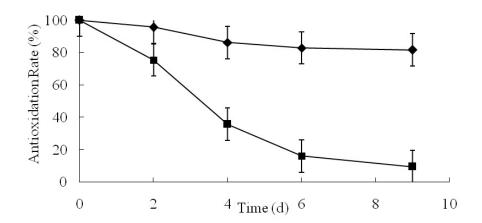


Figure 3. The antioxidative plot of angelica essential oil and microcapsules (■ angelica essential oil ♦ microcapsules).

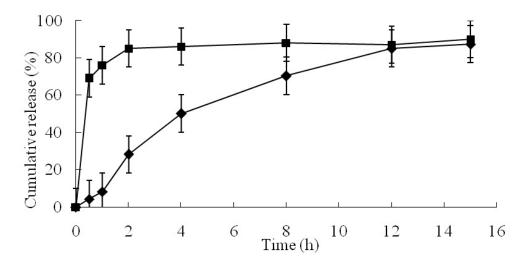


Figure 4. The release profile of angelica essential oil and microcapsules (pH=6.20, gelatin concentration=25.00% (w/w), and core/wall. ratio=40.40%). (■ angelica essential oil ◆ microcapsules)

description of the release profile. P_1 was used to describe the initial phase of drug release and detect any dose burst effect, while t_{85} could ensure that most of the drug was released in a period of time compared to the gastrointestinal residence time (44).

P₁ was significantly influenced by all three formulation variables in a linear manner. Increasing pH at complex coacervation increased P₁. As gelatin concentration or core/wall ratio increased, P₁ significantly decreased. This could be related to an opposite, but parallel effect of these factors on AR, because some studies indicated that the looser surface of microcapsules could result in a marked burst effect and a higher P₁ value (9,43).

 t_{85} was good fitting to quadratic model. It was clear that t_{85} was highly dependent on gelatin concentration. As gelatin concentration increased, the surface of microcapsules was more compact, and t_{85} increased significantly. This corresponded to the higher, positive coefficient of X_2 . This was consisted with the related results of AR and P_1 . However, the negative coefficient of X_3^2 indicated a decrease in the t_{85} at extremely high core/wall ratios.

The *in-vitro* drug release behavior of the newly developed GCM containing AO showed prolonged AO release over 12 h (Figure. 4).

The P₁ of GCM was only 8.29%, but P₁ of AO reached 76%. This could be attributed to the fact that gelatin and chitosan formed coacervate, which could reduce the release of AO from the inner phase to the outer phase of the GCM (45, 46). Some studies indicated that the particulate hydrogels, based on covalently crosslinked chitosan and gelatin by glutaraldehyde, could be considered good candidates for drug delivery (47, 48). The strong networks formed by the double crosslinking of the two polymers in sustained or controlled release system could prevent a rapid disintegration during initial stage (49).

The determination coefficient (r²) values of the zero-order, first-order, and Korsmeyer-Peppas release models were 0.9019, 0.6772, and 0.9486, respectively. The well known Higuchi model ($r^2 = 0.9767$, higher than others) was found to adequately describe the entire release profile (i.e., up to 87% cumulative release) over time. This equation had been used previously to describe release from systems similar to those studied here (21). In addition, Korsmeyer-Peppas model was found to be closer to the best-fit Higuchi model. The values of release exponent (n) determined from in-vitro drug release data was 1.08, indicating the drug release from these microcapsules followed by the super case-II transport mechanism controlled by swelling

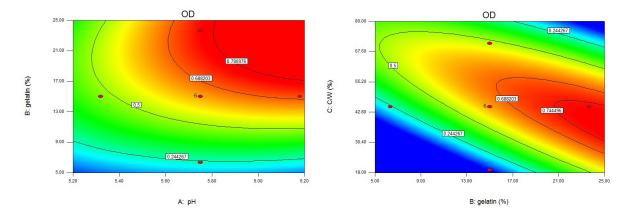


Figure 5. Contour plots of overall desirability as a function of three independent variables. a) gelatin concentration vs. pH with core/wall ratio = 45%; b) core/wall ratio vs. gelatin concentration with pH = 5.75. C/W = core/wall ratio.

and relaxation of polymeric matrix. This could be attributed due to polymer dissolution and enlargement or relaxation of polymeric chain (50).

OD and Prediction

The use of an OD function allowed prediction of the ranges of independent variables where the preferable formulation(s) could occur. The OD was dependent on all the investigated independent variables. This could be seen in the fitting equation for OD (Table 3), which represented the relations between OD and the independent variables. pH at complex coacervation had a positive impact on the OD (positive coefficient of X₁). This could be attributed to the increase in yield, EE, P₁, and t₈₅. However, increasing pH decreased AR. The positive effects of gelatin concentration and core/wall ratio on the OD (positive coefficients of X_2 and X_3) were the results of their positive effects on yield, EE, AR, and t₈₅. However, as gelatin concentration or core/wall ratio increased, P, decreased. Studies had shown that the optimal factor settings for one performance characteristic were not necessarily compatible with those of other performance characteristics. In more general situations, finding compromising conditions on the input variables might be considered that were somewhat favorable to all responses (51, 52). It could be seen from Figure 5. and Table 3. that the highest OD could be achieved at high

pH, high gelatin concentration, and intermediate core/wall ratio. The optimized formulation was prepared at pH 6.20, gelatin concentration 25.00%, and core/wall ratio 40.40%. The measured responses (yield 88.51%, EE 94.85%, AR 81.77%, P_1 8.29%, and t_{85} 15.93 h) of that formulation were very close to the predicted values, and all deviations were less 3%.

Both of two plots (Figure 5) showed that the surface to be relatively flat near the maximum, meaning that small departures from optimality of the variable values, would not appreciably decrease the responses (29).

Conclusion

An AO-loaded microcapsule formulation with excellent antioxidation and sustained-release properties was optimized using RSM by fitting a second-order model to the response data. The model was found to be satisfactory for describing the relationships between independent variables and individual response variables, as well as the relationships between independent variables and the OD. The use of a desirability function appeared to be a useful approach for handling the problem of multiple responses in this case. The optimization method enabled us to predict the values of response variables and OD within the experimental range with good agreement between the predicted and experimental values. An optimum desirability of GCM was achieved

at pH 6.20, gelatin concentration 25.00%, and core/wall ratio 40.40%.

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References

- Hook IL. Danggui to Angelica sinensis root: are potential benefits to European women lost in translation? A review. J. Ethnopharmacol. (2014) 152: 1–13.
- (2) Sa RN, Wang YL, Zhu SQ, Pan XB, Yang YW and Sun YJ. Study on the dynamic variations of Z-ligustilide and n-butylidenephthalide content in essential oil of Radix Angelicae sinensis from different growth period. Zhong Yao Cai. (2012) 35: 1738-42.
- (3) Wedge DE, Klun JA, Tabanca N, Demirci BD, Ozek T, Baser KH, Liu Z, Zhang S, Cantrell CL and Zhang J. Bioactivity-guided fractionation and GC/MS fingerprinting of *Angelica sinensis* and *Angelica archangelica* root components for antifungal and mosquito deterrent activity. *J. Agric. Food Chem.* (2009) 57: 464-70.
- (4) Peng B, Zhao P, Lu YP, Chen MM, Sun H, Wu XM and Zhu L. Z-ligustilide activates the Nrf2/HO-1 pathway and protects against cerebral ischemia-reperfusion injury in-vivo and in-vitro. Brain Res. (2013) 1520: 168–77.
- (5) Wu XM, Qian ZM, Zhu L, Du F, Yung WH, Gong Q and Ke Y. Neuroprotective effect of ligustilide against ischaemia-reperfusion injury via up-regulation of erythropoietin and down-regulation of RTP801. *Br. J. Pharmacol.* (2011) 164: 332-43.
- (6) Feng Z, Lu Y, Wu X, Zhao P, Li J, Peng B, Qian Z and Zhu L. Ligustilide alleviates brain damage and improves cognitive function in rats of chronic cerebral hypoperfusion. *J. Ethnopharmacol.* (2012) 144: 313-21.
- (7) Doneanu C and Anitescu G. Supercritical carbon dioxide extraction of angelica archangelica L. root oil. J. Supercrit. Fluid. (1998) 12: 59-67.
- (8) Yang JY, Chen HH, Wu J, Gong SX, Chen CQ, Zhang TJ and Wang MJ. Advances in studies on pharmacological functions of ligustilide and their mechanisms. *Chinese Herbal Medicines*. (2012) 4: 26-32.
- (9) Neubauer MP, Poehlmann M and Fery A. Microcapsule mechanics: From stability to function. *Adv. Colloid. Interfac.* (2014) 207: 65-80.
- (10) Dong Z, Ma Y, Hayat K, Jia C, Xia S and Zhang X. Morphology and release profile of microcapsules encapsulating peppermint oil by complex coacervation. *J. Food. Eng.* (2011) 104: 455-60.
- (11) Chang CP, Leung TK, Lin SM and Hsu CC. Release

- properties on gelatin-gum arabic microcapsules containing camphor oil with added polystyrene. *Colloid. Surface. B.* (2006) 50: 136-40.
- (12) Fang Z and Bhandari B. Encapsulation of polyphenols a review. *Trends. Food. Sci. Tech.* (2010) 21: 510-23.
- (13) Anitha A, Deepa N, Chennazhi KP, Nair SV, Tamura H and Jayakumar R. Development of mucoadhesive thiolated chitosan nanoparticles for biomedical applications. *Carbohyd. Polym.* (2011) 83: 66-73.
- (14) Vasconcellos FC, Goulart GAS and Beppu MM. Production and characterization of chitosan microparticles containing papain for controlled release applications. *Powder. Technol.* (2011) 205: 65-70.
- (15) Esposito E, Cortesi R and Nastruzzi C. Gelatin microspheres: influence of preparation parameters and thermal treatment on chemico-physical and biopharmaceutical properties. *Biomaterials*. (1996) 17: 2009-20.
- (16) Nagpal K, Singh SK and Mishra DN. Chitosan nanoparticles: a promising system in novel drug delivery. *Chem. Pharm. Bull.* (2010) 58: 1423-30.
- (17) Maisuthisakul P and Gordon MH. Influence of polysaccharides and storage during processing on the properties of mango seed kernel extract (microencapsulation). Food. Chem. (2012) 134: 1453-60
- (18) Lee YK, Ahn SI and Kwak HS. Optimizing microencapsulation of peanut sprout extract by response surface methodology. Food. Hydrocolloid. (2013) 30: 307-14.
- (19) Mondal B, Srivastava VC, Kushwaha JP, Bhatnagar R, Singh S and Mall ID. Parametric and multiple response optimization for the electrochemical treatment of textile printing dye-bath effluent. *Sep. Purif. Technol.* (2013) 109: 135-43.
- (20) Hassan EE, Parish RC and Gallo JM. Optimized formulation of magnetic chitosan microspheres containing the anticancer agent, oxantrazole. *Pharm. Res.* (1992) 9: 390-7.
- (21) Bose A, Wong TW and Singh N. Formulation development and optimization of sustained release matrix tablet of itopride HCl by response surface methodology and its evaluation of release kinetics. *Saudi. Pharm. J.* (2013) 21: 201-13.
- (22) Lee JS, Chung D and Lee HG. Optimization of calcium pectinate gel beads for sustained-release of catechin using response surface methodology. *Int. J. Biol. Macromol.* (2008) 42: 340-7.
- (23) Silva MC and Andrade CT. Evaluating conditions for the formation of chitosan/gelatin microparticles. *Polímeros: Ciência e Tecnologia.* (2009) 19: 133-7.
- (24) Olmedo RH, Asensio CM and Grosso NR. Thermal stability and antioxidant activity of essential oils from aromatic plants farmed in Argentina. *Ind. Crop. Prod.* (2015) 69: 21-8.
- (25) Chong YM, Chang SK, Sia WCM, Yim HS. Antioxidant efficacy of mangosteen (Garcinia mangostana Linn.) peel extracts in sunflower oil during accelerated storage. *Food Bioscience* (2015) 12: 18-25.

- (26) Calija B, Cekić N, Savić S, Daniels R, Marković B and Milić J. pH-sensitive microparticles for oral drug delivery based on alginate/oligochitosan/Eudragit[®] L100-55 "sandwich" polyelectrolyte complex. *Colloid. Surface. B.* (2013) 110: 395-402.
- (27) Costa P and Sousa Lobo JM. Modeling and comparison of dissolution profiles. Eur. J. Pharm. Sci. (2001) 13: 123-33.
- (28) Pal D and Nayak AK. Development, optimization, and anti-diabetic activity of gliclazide-loaded alginate-methyl cellulose mucoadhesive microcapsules. *AAPS PharmSciTech.* (2011) 12: 1431-41.
- (29) Derringer G and Suich R. Simultaneous optimization of several response variables. J. Qual. Technol. (1980) 12: 214-9.
- (30) Cai M, Wang S and Liang H. Optimization of ultrasound-assisted ultrafiltration of radix astragalus extracts with hollow fiber membrane using response surface methodology. Sep. Purif. Technol. (2012) 100: 74-81.
- (31) Abu-Izza K, Garcia-Contreras L and Lu DR. Preparation and evaluation of zidovudine-loaded sustained-release microspheres. 2. optimization of multiple response variables. J. Pharm. Sci. (1996) 85: 572-6.
- (32) Yilmaz E and Bengisu M. Preparation and characterization of physical gels and beads from chitin solutions. *Carbohyd. Polym.* (2003) 54: 479-88.
- (33) Chung EY, Kim HM, Lee GH, Kwak BK, Jung JS, Kuh HJ and Lee J. Design of deformable chitosan microspheres loaded with superparamagnetic iron oxide nanoparticles for embolotherapy detectable by magnetic resonance imaging. *Carbohyd. Polym.* (2012) 90: 1725-31.
- (34) Sun C, Feng YQ, Zhang B, Li XG, Shao JZ, Han JJ and Chen X. Preparation and application of microcapsuleencapsulated color electrophortic fluid in Isopar M system for electrophoretic display. *Opt. Mater.* (2013) 35: 1410-7.
- (35) Shukla RK and Tiwari A. Carbohydrate polymers: applications and recent advances in delivering drugs to the colon. *Carbohyd. Polym.* (2012) 88: 399-416.
- (36) Ramasamy T, Tran TH, Choi JY, Cho HJ, Kim JH, Yong CS, Choi HG and Kim JO. Layer-by-layer coated lipid–polymer hybrid nanoparticles designed for use in anticancer drug delivery. *Carbohyd. Polym.* (2014) 102: 653-61.
- (37) Abreu FOMS, Oliveira EF, Paula HCB and de Paula RCM. Chitosan/cashew gum nanogels for essential oil encapsulation. *Carbohyd. Polym.* (2012) 89: 1277-82.
- (38) Yuen CWM, Yip J, Liu L, Cheuk K, Kan CW, Cheung HC and Cheng SY. Chitosan microcapsules loaded with either miconazole nitrate or clotrimazole, prepared via emulsion technique. *Carbohyd. Polym.* (2012) 89: 795-801.
- (39) Woranuch S and Yoksan R. Eugenol-loaded chitosan nanoparticles: I. Thermal stability improvement of

- eugenol through encapsulation. *Carbohyd. Polym.* (2013) 96: 578-85.
- (40) Qv XY, Zeng ZP and Jiang JG. Preparation of lutein microencapsulation by complex coacervation method and its physicochemical properties and stability. *Food. Hydrocolloid.* (2011) 25: 1596-603.
- (41) Karimi F, Taheri Qazvini N and Namivandi-Zangeneh R. Fish gelatin/laponite biohybrid elastic coacervates: a complexation kinetics-structure relationship study. *Int. J. Biol. Macromol.* (2013) 61: 102-13.
- (42) Doublier JL, Garnier C, Renard D and Sanchez C. Protein-polysaccharide interactions. *Curr. Opin. Colloid. In.* (2000) 5: 202-14.
- (43) Zhang K, Zhang H, Hu X, Bao S and Huang H. Synthesis and release studies of microalgal oil-containing microcapsules prepared by complex coacervation. *Colloid. Surface. B.* (2012) 89: 61-6.
- (44) Hamed E and Sakr A. Application of multiple response optimization technique to extended release formulations design. *J. Control. Release.* (2001) 73: 329-38.
- (45) Patel KD, Singh RK, Lee EJ, Han CM, Won JE, Knowles JC and Kim HW. Tailoring solubility and drug release from electrophoretic deposited chitosangelatin films on titanium. Surf. Coat. Tech. (2014) 242: 232-6
- (46) Cui L, Jia J, Guo Y, Liu Y and Zhu P. Preparation and characterization of IPN hydrogels composed of chitosan and gelatin cross-linked by genipin. *Carbohyd. Polym.* (2014) 99: 31-8.
- (47) Ocak B. Complex coacervation of collagen hydrolysate extracted from leather solid wastes and chitosan for controlled release of lavender oil. *J. Environ. Manage*. (2012) 100: 22-8.
- (48) Basu SK, Kavitha K and Rupeshkumar M. Evaluation of ionotropic cross-linked chitosan/gelatin B microspheres of tramadol hydrochloride. *AAPS PharmSciTech.* (2011) 12: 28-34.
- (49) Jătariu Cadinoiu AN, Holban MN, Peptu CA, Sava A, Costuleanu M and Popa M. Double crosslinked interpenetrated network in nanoparticle form for drug targeting—Preparation, characterization and biodistribution studies. *Int. J. Pharmaceut.* (2012) 436: 66-74.
- (50) Nayak AK, Pal D and Das S. Calcium pectinatefenugreek seed mucilage mucoadhesive beads for controlled delivery of metformin HCl. *Carbohyd. Polym.* (2013) 96: 349-57.
- (51) Köksoy O. A nonlinear programming solution to robust multi-response quality problem. *Appl. Math. Comput.* (2008) 196: 603-12.
- (52) Hejazi TH, Bashiri M, Dı'az-Garcı'a JA and Noghondarian K. Optimization of probabilistic multiple response surfaces. *Appl. Math. Model.* (2012) 36: 1275-85.

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